

## ON AN ESSENTIALLY COMPLETE CLASS OF THE ESTIMATORS UNDER CERTAIN RESTRICTIONS ON PARAMETERS

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### Summary

There are some results dealing with statistical problem in the presence of the restrictions on parameters. However, it seems that there are no papers which discuss the complete class in the statistical estimation problem under some restrictions by inequalities so long as I know. Our aim is to show that the class of estimators satisfying the same inequalities as that of parameters is essentially complete.

### 1. Introduction

Suppose that  $X$  is the sample space of the observed random variables.  $\mathcal{A}$  is a  $\sigma$ -field of subsets of  $X$  and  $\{P_\theta; \theta \in \Theta\}$  is a family of probability measures on  $\mathcal{A}$ . Here it is assumed that  $\Theta$  is an open and convex subset of a  $p$ -dimensional Euclidean space and  $\mathcal{C}$  is the Borel  $\sigma$ -field on  $\Theta$ . Let  $p(x; \theta)$  designate the probability density function of  $P_\theta$  with respect to a  $\sigma$ -finite measure  $\mu$  on  $\mathcal{A}$ , for which it is assumed to be  $\mathcal{C}$ -measurable for each  $x \in X$ .

As we are interested in the estimation of a parameter  $\theta$ , we may suppose that the action space  $Y$  is identical to the parameter space  $\Theta$ .

Moreover, assume that the loss function is  $L(\theta, y) = (y - \theta)' M (y - \theta)$  where  $(y - \theta)' = (y_1 - \theta_1, \dots, y_p - \theta_p)$  and  $M$  is a known  $p \times p$  real symmetric positive-definite matrix. In the sequel, we consider the estimation problem in which there is the restricted condition on parameters so that  $\theta$  belongs to the set

$$(1) \quad F = \{\theta \in \Theta: f_i(\theta) \leq c_i, i=1, 2, \dots, k\}$$

where  $k$  is a given positive integer,  $f_i$ 's are convex and continuous real-valued functions of  $\theta$  and  $c_i$ 's are constants.

Since the loss function  $L(\theta, y)$  is convex in  $y$  for each  $\theta$ , we may only consider

the class  $D$  of non-randomized decision functions, i. e.,  $D = \{\delta; \delta: X \rightarrow Y, \text{ measurable}\}$  (see Blackwell-Girshick, p. 294). The risk function of  $\delta \in D$  is defined by

$$(2) \quad r(\theta, \delta) = \int_X L(\theta, \delta(x)) p(x; \theta) \mu(dx).$$

For a given  $\delta \in D$ ,  $r(\theta, \delta)$  is considered a function on  $\Theta$ , which is  $C$ -measurable.

Let  $\Xi$  be the set of prior probability measures on  $(\Theta, C)$ . Then the prior risk of  $\delta$  with respect to  $\xi \in \Xi$  is defined by

$$(3) \quad r(\xi, \delta) = \int_{\Theta} r(\theta, \delta) \xi(d\theta).$$

We denote the class of Bayes decision function by  $B$ .

Let  $\Xi_0 = \{\xi \in \Xi; r(\xi, \delta) < \infty \text{ for some } \delta \in D\}$  and  $B_0 = \{\delta \in D; \delta \text{ is Bayes for some } \xi \in \Xi_0\}$ . A sequence  $\{\delta_n\}$  of randomized decision functions is said to *converge regularly* to a randomized decision function  $\delta$  if for every  $\mu$ -integrable function  $f(x)$  and every bounded continuous  $g(y)$

$$\begin{aligned} \lim_{n \rightarrow \infty} \int_X \int_Y f(x) g(y) \delta_n(dy | x) \mu(dx) \\ = \int_X \int_Y f(x) g(y) \delta(dy | x) \mu(dx). \end{aligned}$$

From LeCam (1955) the closure of  $B$  in the topology of regular convergence is an essentially complete class. If  $\delta_n \in D$  converges regularly to  $\delta \in D$ , we have that  $\delta_n(x)$  converges a. e. to  $\delta(x)$ . Accordingly, the closure  $\bar{B}$  of  $B$  in the topology of convergence a. e. is an essentially complete class.

## 2. Complete class theorems

In order to achieve our aim, one theorem and some lemmas are prepared.

**THEOREM 1.**  $\bar{B}_0$  is essentially complete.

**PROOF.** Let  $B_1$  be the part of  $\bar{B}$  which excludes those  $\delta$ 's with infinite risk. Then  $B_1$  is essentially complete because  $\bar{B}$  is. Let  $\delta \in B_1$ , i. e.,  $\delta \in \bar{B}$  and  $r(\theta, \delta) < \infty$ . Then there is a sequence  $\{\delta_n\}$  of  $B$  such that  $\delta_n$  converges regularly to  $\delta$  and  $r(\theta, \delta_n) < \infty$  for all  $n$ . In fact, for the regular convergence of  $\{\delta_n\}$  to  $\delta$ , we have that  $\lim_{n \rightarrow \infty} r(\theta, \delta_n) = r(\theta, \delta) < \infty$  for every  $\theta$  (see Farrell). Hence there is a  $\xi_n \in \Xi$  such that  $\delta_n$  is Bayes for  $\xi_n$  and  $r(\xi_n, \delta_n) < \infty$ ; that is,  $\xi_n \in \Xi_0$ . This shows that

there exists a sequence  $\{\delta_n\}$  of  $B_0$  such that  $\delta_n$  converges regularly to  $\delta$ ; that is,  $\delta \in B_0$ . Hence  $B_1 \subset \bar{B}_0$ . This is the desirable result because  $B_1$  is essentially complete.

We denote by  $D_F$  the class of all estimators  $\delta$  which satisfy the inequalities  $f_i(\delta) \leq c_i$ ,  $i=1, 2, \dots, k$ .

LEMMA 1.  $B_0 \subset D_F$ .

PROOF. Due to Girshick and Savage,  $\delta \in B_0$  is given by

$$(4) \quad \delta(x) = \int_F \theta \xi(d\theta | x) \text{ a. e.}$$

for some  $\xi \in \mathcal{E}_0$ .

From Jensen's inequality, we have

$$\begin{aligned} f_i(\delta(x)) &= f_i\left(\int_F \theta \xi(d\theta | x)\right) \\ &\leq \int_F f_i(\theta) \xi(d\theta | x) \\ (5) \quad &\leq c_i \end{aligned}$$

for  $i=1, 2, \dots, k$ . This shows *Lemma 1*.

LEMMA 2.  $\bar{D}_F = D_F$ .

PROOF. Let  $\delta(x)$  be any element of  $\bar{D}_F$ . Then there exists a sequence  $\{\delta_n(x)\}$  in  $D_F$  such that  $\delta_n(x)$  converges a. e. to  $\delta(x)$ . Since  $f_i(\delta_n(x)) \leq c_i$  for all  $n$  and  $f_i$  is continuous for  $i=1, 2, \dots, k$ , it is clearly implied that  $f_i(\delta(x)) \leq c_i$  for  $i=1, 2, \dots, k$ .

The following theorem is our aim.

THEOREM 2.  $D_F$  is an essentially complete class.

PROOF. From *Lemma 1*,  $B_0 \subset D_F$ . Hence  $\bar{B}_0 \subset \bar{D}_F$ . Hence, from *Lemma 2*,  $\bar{B}_0 \subset D_F$ . This shows *Theorem 2*.

### 3. Examples

EXAMPLE 1. Let  $X_1, X_2, \dots, X_n$ , be independent and identically distributed random variables with an unknown mean  $\theta$ . When it is known previously that  $a \leq \theta \leq b$  ( $a = -\infty$  or  $b = +\infty$  may be permitted), we consider the estimation

problem of  $\theta$ . Since the restriction  $a \leq \theta \leq b$  can be written as follows: Let  $f_1(\theta) = -\theta$  and  $f_2(\theta) = \theta$ , and let  $c_1 = -a$  and  $c_2 = b$ . Then  $F = [a, b] = \{\theta; f_1(\theta) \leq c_1 \text{ and } f_2(\theta) \leq c_2\}$ .

Since both  $f_1$  and  $f_2$  are convex and continuous in  $\theta \in [a, b]$  we can apply *Theorem 2* to this problem.

We can directly have from

$$\hat{\theta}_\xi(x) = \frac{\int_a^b \theta p(x; \theta) \xi(d\theta)}{\int_a^b p(x; \theta) \xi(d\theta)} \text{ and } a \leq \theta \leq b$$

that  $a \leq \hat{\theta}_\xi(x) \leq b$ .

EXAMPLE 2. Let  $X_1$  and  $X_2$  be independent random variables with means  $\theta_1$  and  $\theta_2$ . Under the restriction  $\theta_1^2 + \theta_2^2 \leq a^2$  ( $a$  is a positive constant), we consider the estimation of  $\theta_1$  and  $\theta_2$ . Such a restriction can be written as follows.

$$\begin{aligned} F &= \{\theta = (\theta_1, \theta_2): \theta_1^2 + \theta_2^2 \leq a^2\} \\ &= \{\theta; f_1(\theta_1, \theta_2) = \theta_2 - \sqrt{a^2 - \theta_1^2} \leq 0 \\ &\quad \text{and } f_2(\theta_1, \theta_2) = -\theta_2 - \sqrt{a^2 - \theta_1^2} \leq 0\} \end{aligned}$$

We can easily check the conditions which  $f_1$  and  $f_2$  must satisfy.

It is directly shown that *Theorem 2* can be applied to this problem. Indeed,

$$\hat{\theta}_1 = \frac{\int_F \theta_1 p(x; \theta) \xi(d\theta)}{\int_F p(x; \theta) \xi(d\theta)} \text{ and } \hat{\theta}_2 = \frac{\int_F \theta_2 p(x; \theta) \xi(d\theta)}{\int_F p(x; \theta) \xi(d\theta)}$$

where  $x = (x_1, x_2)$ .

From Schwarz's inequality, we have

$$\begin{aligned} \hat{\theta}_1^2 + \hat{\theta}_2^2 &= \frac{(\int_F \theta_1 p(x; \theta) \xi(d\theta))^2 + (\int_F \theta_2 p(x; \theta) \xi(d\theta))^2}{(\int_F p(x; \theta) \xi(d\theta))^2} \\ &\leq \frac{\int_F \theta_1^2 p(x; \theta) \xi(d\theta) + \int_F \theta_2^2 p(x; \theta) \xi(d\theta)}{\int_F p(x; \theta) \xi(d\theta)} \\ &= \frac{\int_F (\theta_1^2 + \theta_2^2) p(x; \theta) \xi(d\theta)}{\int_F p(x; \theta) \xi(d\theta)} \\ &\leq a^2. \end{aligned}$$



EXAMPLE 3. Let  $X_1$  and  $X_2$  be for as *Example 2*. We suppose that there is an order restriction  $a \leq \theta_1 \leq \theta_2 \leq b$  ( $a = -\infty$  or  $b = +\infty$  may be permitted). This restriction can be expressed using functions  $f_1(\theta_1, \theta_2) = -\theta_1$ ,  $f_2(\theta_1, \theta_2) = \theta_1 - \theta_2$  and  $f_3(\theta_1, \theta_2) = \theta_2$ . That is,

$$\begin{aligned} F &= \{\theta = (\theta_1, \theta_2); a \leq \theta_1 \leq \theta_2 \leq b\} \\ &= \{\theta; f_1(\theta) \leq -a, f_2(\theta) \leq 0 \text{ and } f_3(\theta) \leq b\}. \end{aligned}$$

It is easily seen that  $f_1$ ,  $f_2$  and  $f_3$  satisfy the conditions.

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### References

- [1] Girshick, M. A. and Savage, L. J. (1951). Bayes and minimax estimators for quadratic loss functions. *Proc. Second Berkeley Symp. Math. Statist. Prob.* 1 53-74.
- [2] Blackwell, D. and Girshick, M. A. (1954). *Theory of Games and Statistical Decisions*. Wiley.
- [3] LeCam, L. (1955). An extension of Wald's theory of statistical decision functions. *Ann. Math. Statist.* 26 69-81.
- [4] Sacks, J. (1963). Generalized Bayes solutions in estimation problems. *Ann. Math. Statist.* 34 751-768.
- [5] DeGroot, M. H. and Rao, M. M. (1963). Bayes estimation with convex loss. *Ann. Math. Statist.* 34 839-846.
- [6] Farrell, R. H. (1967). Weak limits of sequences of Bayes procedures in estimation theory. *Fifth Berkeley Symp. Math. Statist. Prob.* 1 83-111.
- [7] Zacks, S. (1971). *The Theory of Statistical Inference*. Wiley.
- [8] Barlow, R. E., Bartholomev, D. J., Bermner, J. M. and Brunk, H. D. (1972). *Statistical Inference under Order Restrictions*. Wiley.

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